*Article*

# Financial Year and Calendar Year Effects Across Indian Sectors: A Revisit from Investors Perspective

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**Parul Bhatia[1] and Niyati Chaudhary[2]**

## Abstract

Anomalies in stock markets have been studied with varied approaches worldwide, and mixed outcomes exist. The present study has investigated the financial year and calendar year effect for indices from Indian markets. Sectoral indices from Indian markets have been incorporated in this study on the basis of their weightage in the stock market. Data from 2011 to 2023 have been engaged with historical prices of these indices taken from the stock exchange websites. The analysis has been carried out using the ordinary least squares regression method and the independent sample *t*-test. Weak anomalies have been documented in the case of the financial services sector. Therefore, the findings indicate rare possibilities for investors to gauge abnormal returns with strategies related to anomalies. With increasing transparency in the online trading mechanism, such possibilities cease to exist for investors. Moreover, market dynamics have transformed on account of several shocks such as COVID-19, Omicron, Russia–Ukraine, Visa-related international issues in the world over economies and financial markets. Investors and markets have become more cautious and look out for a precautionary approach while dealing in stock market trading.

## Keywords

Financial year, calendar year, OLS, stock returns, investor, anomalies

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[1] Shri Vishwakarma Skill University, Dudhola, Palwal, Haryana, India
[2] SGT University, Gurugram, Haryana, India

**Corresponding Author:**
Parul Bhatia, Shri Vishwakarma Skill University, Dudhola, Palwal, Haryana 121102, India.
E-mail: parul.bhatia84@gmail.com

## Introduction

Anomalies prevail at varied intervals of time and may be of different kinds in the stock markets (El et al., 2021) like day, week or month. Stock returns may be high or low due to such anomalies offering opportunities at various stages to investors. The efficient market hypothesis EMH— also known as efficient market theory— provides theoretical variant in the pricing of stocks, which explains the concept of market efficiency (Neeraja & Srikanth, 2014). Adaptive market hypothesis (AMH) is a more refined variant that indicates the fluctuation in the market and possible factors (Ito & Sugiyama, 2009; Kim et al., 2011). Several studies in the field of anomalies have provided mixed outcomes in Indian and international stock markets (Hiremath & Kumari, 2014; Lim et al., 2008; Noda, 2012; Rosini & Shenai, 2020; Urquhart & Hudson, 2013).

The results from Indian stock markets have reflected different anomalies in diverse periods in line with EMH theory. It has been shown that anomalies and their effect vary over market conditions as well as investor expectations (Hiremath & Kumari, 2014; Shahid & Sattar, 2017). Efficient stock markets narrate that all information stands available affecting the prices meaning that investors would have no chance to earn abnormal profits (Al-Rjoub & Alwaked, 2010). This phenomenon has been further studied with new available information in the market which may be more appropriate in return prediction in comparison to past prices of the stock (Chaker & Sabah, 2018). Further, Urquhart and McGroarty (2014) mentioned that there may be no such condition of efficient market rather continuous variation in market due to the seasonal anomalies. The identification of anomalies sparks investigations into the information efficiency of stock markets, providing a platform for market participants to craft successful trading strategies and potentially secure returns surpassing the typical benchmarks. Over the past four decades, changes in calendar effects have been a focal point of research in African stock markets, commodity markets to name a few (Andrew & Seth, 2023; Chhabra & Gupta, 2022; El et al., 2021; Li et al., 2023). Anomaly may be an important element to be analysed for the estimation of stock returns (Guo et al., 2020). It depends upon the risk perception of the investors and uncertainty in the markets (Chaudhary et al., 2020; Engelberg et al., 2020).

Indian stock market has experienced more volatility during the COVID-19 period. However, market efficiency theory mentioned that if markets are rational, then prices reflect all available information (Dash & Mahakud, 2015). The calendar year effect vanished over the time (Sawitri & Astuty, 2018), and calendar anomalies diminished post the global financial crisis period in various sectoral indices in the Bursa Malaysia stock exchange (Kaur et al., 2019). Further events such as holidays, climate changes and policy regime alterations may affect the returns of various stocks. Anomalies pertain globally, but the January effect has been one of the most important anomalies. In addition, the calendar effect has been an important anomaly affecting the returns of stocks (Floros & Salvador, 2014). These effects have been explored by the agents and potential investors

in financial markets for obtaining abnormal returns (Chaudhary et al., 2020). Indian markets have been more volatile than the global markets. Liquidity, transparency and investment culture have been the major contributors to such variations (Izzeddien, 2021). December and March have also been studied for the Karachi Stock Exchange and the Pakistan Stock Exchange showing inefficiencies (Anjum, 2020; Elangovan, 2022). Also, seasonality deviations have been prevalent in longer periods for such anomalies (Kaushik & Nagarkar, 2023).

**Table 1.** Calendar Year and Financial Year Effect Studies in Indian Markets.

| Authors (Year) | Sample | Result |
|---|---|---|
| Verma & Kumar (2012) | (Bombay Stock Exchange) BSE- Sensex (Jan 1991–Dec 2020) | The Bombay Stock Exchange does not exhibit any month-of-the-year effect. |
| Debasish (2012) | Specific sectors of Indian market (gas, oil and refineries) (Jan 2006–Dec 2010) | A month-of-the-year effect has been observed, primarily occurring in September, August or February. |
| Deepak & Viswanath (2012) | (National Stock Exchange) NSE NIFTY (monthly data) (Jan 1990–Dec 2011) | Effects of March, May and October months found in this study. |
| Pathak (2013) | S&P CNX NIFTY (NSE) (Apr 2007–Mar 2012) | The absence of the month-of-year effect indicates the lack of seasonality in the Indian stock market. |
| Sriram & Devi (2013) | BSE Sensex (Apr 2004–Mar 2012) | The findings do not validate the presence of seasonality in stock returns, including the January effect. |
| Yadav (2013) | NSE (Jan 1996–Dec 2013) | There is a discernible month-of-the-year effect observed in both market volatility and market returns. |
| Nageswari et al. (2013) | S&P CNX NIFTY (NSE) and S&P CNX 500 (Apr 2002–Mar 2011) | The data did not reveal a January anomaly during the study period. |
| Neeraja & Srikanth (2014) | BSE (IT) and BSE (Sensex) Jan 1999–Dec 2013 | The data suggest a seasonal pattern in the Indian IT sector. |
| Patel (2014) | NSE-CNX 500 Index (June 1999–June 2012) | There is no single month that consistently predicts the performance of the Indian stock market. |
| Lodha & Soral (2015) | BSE (major indices) | The months of September and December saw significant returns. |

*(Table 1 continued)*

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| Authors (Year) | Sample | Result |
|---|---|---|
| Sudarvel & Velmurugan (2015) | BSE (bank index) (Jan 2002–June 2015) | January effect was present during the study period. |
| Choithala & Ajmal (2016) | BSE-Sensex (June 1999–June 2012) | Indian stock returns exhibit a significant December effect in volatility. |
| Kumar & Dawar (2017) | BSE (major indices) (Jan 1999–Dec 2015) | Calendar effect was absent in this index during this period. |
| Kumari & Uthra (2018) | BSE (healthcare index) (Apr 2012–Mar 2016) | While the turn-of-the-month effect remains, its occurrence has exhibited a recent shift towards earlier dates. |
| Jain (2019) | Gold and crude oil market in India | April effect was found in this study, instead of January effect. |
| Singh & Das (2020) | BSE (Bank & IT indices) (2010–2019) | January effect is present in these indices. |
| Bhatia (2021) | BSE Index (1991–2019) | No anomalies found in the Indian stock market. |
| Bankoti (2021) | BSE and NSE (Apr 2010–Mar 2019) | Calendar anomalies were present for these specific period. |
| Elangovan et al., 2022 | NIFTY (500) and S&P BSE 500 | March was effective in the month-of-the-year effect among other months. |

## Contribution of the Study

This study has examined sectoral indices within the Indian stock market laying emphasis on identified anomalies. The time period selected for the study and its impact on anomalies have not been explored before for investors' understanding. Thus, the present work shall be a guide for investors to make informed decisions and understand the complexity and behaviour of markets during possible shocks. Calendar and financial year effects, which have been earlier recognised in the studies, have not been present significantly in the results. Thus, a newer direction and strategy in improved mechanisms of stock market trading may be formulated with the help of results of the study. Additionally, by examining anomalies over an extended timeframe, the study identifies more robust and persistent trends, offering greater confidence to investors when making investment decisions. The result of the present study will also add to the existing literature and will prove helpful for the investors to take correction actions while investing in the Indian stocks. The twin effect of the financial year and calendar year assists the academicians and potential investors to find out the better return for the stocks in various sectors.

The manuscript has been divided into six sections. The first section explains the introduction. The second section narrates the previous studies existing in the literature. The research methodology has been explained in the third section. The fourth section presents the analysis and findings. The fifth section concludes

the study, and the sixth section reflects upon the further scope in the direction of anomaly studies.

## Literature Review

### *Calendar Year Effects*

Stock market seasonality is known as calendar anomalies (Alvarado & Demmler, 2019). The January month is also having the tax hypothesis effect. The January effect is one of the most popular effects of calendar anomalies in which investors expect a higher return than other months. Andrew and Seth (2023) studied the calendar anomalies in petroleum and petroleum products and found variation in the return across the energy commodities. Few authors found the January effect in their study (Gouider et al., 2015; Selvakumar, 2011; Singh & Das, 2020) and specified the reason that investors buy at the starting of the year with higher return expectations. Rosini and Shenai (2020) mentioned the time-variant behaviour of calendar anomalies in the US stock market (Faisal & Majid; 2016). For example, in December, investors sell their shares for tax loss, leading to a decline in the prices of shares, whereas by the end of December, they purchase shares that increases the prices of shares and has a positive impact on the return of stocks in January. For different stock markets, different results were found by many researchers. Researchers such as Floros (2008) (Athens), Ariss et al. (2011) (Gulf Cooperation Council indices), Iqbal et al. (2013) (Pakistan) and Al-Saad & Moosa (2005) (Kuwait), Kaur et al. (2019) (Malaysia) found no January effect on return of stocks, whereas Wong et al. (2006) found positive and significant return for the Singapore stock market. There was no January effect in BSE and NIFTY returns (Kaur, 2004), although the February and April effects were observed in NIFTY instead of the January effect (Jain, 2019; Saxena et al., 2021). No calendar anomalies in France, Germany, Italy and Spain stock exchange indices were found (Rossi & Gunardi, 2018). The August and November effects in Bursa Malaysia stock exchange were observed (Kaur et al., 2019). Acharya et al. (2022) mentioned about the September effect in the Indian stock market, because in India, the festive season starts from September.

Different month effects are found in different countries, such as the December effect is found in India by Parikh (2009), while the May effect is found by Purohit and Tyagi (2015) and the positive November effects are found by Chakrabarti and Sen (2008) for the firms listed in the Shanghai Stock Exchange. In the Karachi Stock Exchange, high return is observed in December, while in the Pakistan Stock Exchange, high return is noted in March (Anjum, 2020). The year-end effect or income tax savings or festive season may be the reason for such results (Faisal & Majid, 2016; Purohit & Tyagi, 2015; Sarbapriya, 2012). While the September month had a positive effect on the Iranian capital market because of Eid (Ansari & Jafari, 2020). The September effect was found in the Indonesian Stock Exchange, the May effect was found in Spain's IBEX35, and the February effect was found in England's FTSE100 (Sawitri & Astuty, 2018).

## *Financial Year Effects*

Negative return in the month of March was found by Sarbapriya (2012). Some researchers stated the absence of turn-of-the-month effect in the return of stocks in the Indian stock market (Kushwah & Munshi, 2018; Nageswari et al., 2011) and in the Mexican stock market (Alvarado & Demmler, 2019; Ansari, 2020). No anomalies in the Indian stock market, especially in March and April months, were observed for stocks (Bhatia, 2021), while Elangovan (2022) observed a month-of-the-year effect present in March in Indian stock exchange.

Qadan et al. (2021) studied the seasonal and calendar anomalies in cryptocurrencies for measuring the pricing efficiency. Rare opportunities are there in Indian stock market for higher return for investors (Bhatia, 2021). Li et al. (2024) and Kaushik and Nagarkar (2023) also stated the presence of seasonal variations in long tenure data.

## *Sectoral Effects*

Chaudhary et al. (2020) revealed that during the Covid period, the banking sector, the realty sector and capital goods had the poorest return, while the healthcare sector and the IT sector showed a positive return. Inefficient information is the main source of anomalies in different sectors of the Indonesian stock market (Said et al., 2018). Covid had a less negative impact on anomalies in the Indian stock market than its peers (Sudha et al., 2020). Although the effect of seasonal anomalies disappeared in the post-crisis period (Adam et al., 2016), the calendar effect was found in different sectors of the Mexican stock market, such as the health care sector, financial sector, consumer sector and telecommunication sector (Alvarado & Demmler, 2019). Cristi et al. (2020) suggested to invest in the low volatility stocks of pharmaceutical, fast moving consumer goods (FMCG) and IT sector in the Indian stock market when preparing their investment strategy for high return. While the IT sector of the Indian stock market was found to be a weak form of efficiency, the January and turn-of-month effects on return and further volatility were mentioned in the IT and banking sectors of India (Singh & Das, 2020). Multi-commodity exchange was studied by Chabra and Gupta (2022), and they found the presence of the anomalies effect in the Indian stock exchange. Li et al. (2022) studied the crude oil market and reported the same. While the absence of anomalies was reported by Bankoti (2021), Li et al. (2023) and Li et al. (2022) mentioned that the effect of anomalies declined with high volatility in the US stock market.

## *Research Gap*

The existing literature on anomalies in the Indian stock markets has shown the presence of the mixed form of anomalies. These inefficiencies persist in the form of day/week/month arrangements where studies indicate the presence/absence of anomalies in various industries and stocks (Engle et al., 1990; Tadepalli & Jain, 2018). Calendar year and financial year anomalies have also been tested earlier with varied research methods for stock market indices and individual stocks

(Ansari, 2020; Bhatia, 2021; Jain, 2019; Kushwah & Munshi, 2018; Singh & Das, 2020). However, there remains a strong gap in analysing the sectoral indices for the time period considered in the present study. Therefore, with this strong notion based on the current literature, the present work will contribute to the understanding of the investors. Attention may be paid to the existing regime of stock market anomalies before plugging the funds into the market.

### *Hypotheses*

$H_1$: There exists a calendar year effect for sectoral indices in Indian markets.
$H_2$: There exists a financial year effect for sectoral indices in Indian markets.
$H_3$: There is a spillover from March returns to April returns for Indian sectoral indices.
$H_4$: There is a spillover from December returns to January returns for Indian sectoral indices.

## Research Methodology

### *Data*

The historical indices for automobiles (Auto), financial services (FS), media, FMCG, pharmaceuticals (Pharma), metal, realty, information technology (IT), energy and banking (Bank) have been incorporated. These historical values of 2011–2019 have been downloaded from the website of the National Stock Exchange. Due to Covid impact, data after December 2019 have been avoided as the results may not be reliable during this period. The historical returns have been computed with the help of the following formula:

$$\text{Rit} = P_t - (P_t - 1) / (P_t - 1) * 100. \quad (1)$$

### *Variables*

The objectives of the study required monthly data of four months, in particular for Indian sectoral indices. The investigation of the financial year effect called for March and April returns, while the calendar year effect desired December and January returns. Companies may make major announcements during these months due to the closing of the books of accounts. Therefore, these four months have been identified and incorporated in the study to explore anomalies across different sectors in the Indian economy.

### *Sample Selection*

Indian markets may be split into a variety of sectors such as banking, IT, automobiles and pharmaceuticals. The sector-wise contribution to stock markets as well as Indian economy however remains prevalent by a few sectors. The study

**Table 2.** Indices in the Study.

| Variable | Source | Justification |
|---|---|---|
| Auto | NSE website | Benchmark index for the automobile sector consisting of 15 stocks. The companies included range over cars, motorcycles, auto ancillaries, tyres and heavy vehicles. |
| Bank | NSE website | Resembles 12 Indian banking stocks based on their liquidity and market capitalisation. |
| Energy | NSE website | One of the most significant drivers for growth of Indian economy. Considered as a crucial input for sustainable development. Index represents the performance of such companies. |
| FMCG | NSE website | Consisting of 15 stocks of FMCG companies. Such companies are involved with the production of non-durable goods the can be used for mass consumption. |
| Financial services | NSE website | Demonstrates the capital market performance of 15 financial service stocks. These include insurance companies, housing banks, banks, financial institutions in the list. |
| IT | NSE website | Base index for the performance of companies which are involved with IT and IT-enabled services. The IT segment of the Indian market is well resembled with this index. |
| Media | NSE website | This index includes companies' stocks that are involved in media, entertainment, printing and publishing. Captures the behaviour and performance of such companies. |
| Pharma | NSE website | This index captures the performance of pharmaceutical companies from Indian markets. This sector has been one of the key drivers towards growth of the Indian economy after software companies. |
| Realty | NSE website | Public funds and private equity both have witnessed a tremendous growth after real estate companies entering the stock markets. This index has been incorporated with a viewpoint that it carries immense potential for growth of investors. |
| Metal | NSE website | Reflects the performance of 15 stocks listed on NSE. The free float market capitalisation remains the base for its index values. |

has tried to incorporate such sectors for investigating the presence of anomalies across Indian markets. A brief justification for sample selection has been provided in Table 2.

## *Test for Stationary Series*

Time series analysis and modelling require the data to be stationary at the outset. The preliminary assumption of stationary series has been tested with the augmented Dickey–Fuller statistic. The index returns have been found stationary at level and thus integrated at $I(0)$.

### Independent Sample t-Test

Initially, the significant difference between two groups of returns, namely March–April and December–January, have been tested. This investigation has been carried out with the independent sample *t*-test to test difference between March–April and December–January returns.

### Model Specification

#### Ordinary Least Squares

The cause-and-effect relationship for March–April and December–January has been investigated using the following equation:

$$C_t + x10t + et. \tag{2}$$

In the first set-up, causation from March to April has been investigated, and in the second one, December to January has been examined. The linear relation between independent and dependent variables has been established with the ordinary least squares (OLS) model.

## Analysis and Findings

### Descriptive Statistics

The investigation of hypotheses, descriptive statistics and other preliminary assumptions for fitting the models have been detailed in this section. Table 3 shows the descriptive statistics for all 10 sectors' indices incorporated in the study.

The mean values, standard deviation, and maximum and minimum values during the four months (March, April, December and January) have been presented. The mean average returns and standard deviations have been closely monitored anticipating results from companies in these sectors in the months of March and December. Thereby, the impact of such information from management of companies may be seen in the months of April and January respectively. It has been observed that in the majority of comparisons between March and April returns, negative returns have been presented in April. The exceptions however have been media, metal, pharmaceuticals and IT. The examination further with OLS has been done to explore these observations. Also, standard deviation in few cases, such as auto, financial services, FMCG, media, realty, energy and ban, has been higher in December than in Januaryk. This does not correlate with the higher returns observed in December than in January. Ruling out the possibility of chance factor, further these variables have been investigated with OLS.

### ADF Results

The preliminary assumption for time series modelling remains the check for the unit root, which has been shown in Table 4.

**Table 3.** Descriptive Statistics.

| Sector | Month | Minimum | Maximum | Mean | Std. Deviation |
|---|---|---|---|---|---|
| Auto | March | –0.261 | 0.396 | 0.098 | 0.228 |
| | April | –0.331 | 0.333 | 0.090 | 0.217 |
| | December | –0.263 | 0.449 | 0.143 | 0.240 |
| | January | –0.346 | 0.578 | 0.109 | 0.298 |
| FS | March | –0.297 | 0.523 | 0.148 | 0.247 |
| | April | –0.172 | 0.387 | 0.147 | 0.191 |
| | December | –0.076 | 0.453 | 0.168 | 0.201 |
| | January | –0.205 | 0.600 | 0.142 | 0.252 |
| Media | March | –0.876 | 0.396 | 0.025 | 0.391 |
| | April | –0.731 | 0.361 | 0.028 | 0.353 |
| | December | –0.353 | 0.462 | 0.044 | 0.272 |
| | January | –0.473 | 0.397 | 0.033 | 0.272 |
| FMCG | March | –0.104 | 0.293 | 0.124 | 0.122 |
| | April | –0.057 | 0.326 | 0.113 | 0.115 |
| | December | –0.013 | 0.396 | 0.134 | 0.130 |
| | January | –0.101 | 0.391 | 0.129 | 0.138 |
| Pharma | March | –0.264 | 0.537 | 0.099 | 0.300 |
| | April | –0.111 | 0.398 | 0.107 | 0.201 |
| | December | –0.153 | 0.474 | 0.115 | 0.229 |
| | January | –0.138 | 0.418 | 0.101 | 0.220 |
| Metal | March | –0.652 | 0.919 | 0.029 | 0.459 |
| | April | –0.507 | 0.958 | 0.054 | 0.436 |
| | December | –0.376 | 0.396 | 0.031 | 0.267 |
| | January | –0.416 | 0.616 | 0.000 | 0.326 |
| Realty | March | –0.428 | 0.644 | 0.037 | 0.344 |
| | April | –0.321 | 0.503 | 0.035 | 0.301 |
| | December | –0.421 | 0.741 | 0.059 | 0.377 |
| | January | –0.646 | 0.652 | 0.031 | 0.447 |
| IT | March | –0.202 | 0.706 | 0.153 | 0.263 |
| | April | –0.169 | 0.598 | 0.160 | 0.257 |
| | December | –0.075 | 0.457 | 0.153 | 0.190 |
| | January | –0.132 | 0.423 | 0.153 | 0.188 |
| Energy | March | –0.393 | 0.491 | 0.098 | 0.248 |
| | April | –0.227 | 0.378 | 0.097 | 0.180 |
| | December | –0.007 | 0.327 | 0.099 | 0.106 |
| | January | –0.121 | 0.303 | 0.080 | 0.135 |
| Bank | March | –0.463 | 0.554 | 0.131 | 0.291 |
| | April | –0.324 | 0.420 | 0.129 | 0.232 |
| | December | –0.102 | 0.498 | 0.152 | 0.227 |
| | January | –0.246 | 0.662 | 0.125 | 0.285 |

The results have been analysed with intercept, without the presence of any trend for all variables across the 10 sectors. The outcomes in Table 4 reflect that the data have been stationary as the minimum $P$ value is <.05. (Rejection of $H_0$: *The series has a unit root*.). Thus, the OLS model for testing the financial year and calendar year effects has been successfully carried out.

**Table 4.** Unit Root Results.

| Augmented Dickey–Fuller Statistics | | | | | | | | | | | | |
|---|---|---|---|---|---|---|---|---|---|---|---|---|
| $H_0$: The series has a unit root. | | | | | | | | | | | | |
| | March | | | April | | | December | | | January | | |
| Sector | Intercept | Trend | Without Trend | Intercept | Trend | Without Trend | Intercept | Trend | Without Trend | Intercept | Trend | Without Trend |
| Auto | −2.510 | −3.279 | −2.564** | −2.776 | −3.227 | −2.030** | −2.871* | −7.831*** | −2.758** | −3.302* | −5.813** | −3.059*** |
| | (0.1469) | (0.157) | (0.018) | (0.109) | (0.165) | (0.047) | (0.091) | (0.004) | (0.013) | (0.051) | (0.015) | (0.007) |
| FS | −5.689*** | −5.430** | −1.080 | −5.956*** | −5.531** | −0.850 | −6.866*** | −6.240*** | −0.788 | −5.671*** | −5.013** | −3.762*** |
| | (0.0029) | (0.015) | (0.227) | (0.002) | (0.014) | (0.313) | (0.001) | (0.007) | (0.339) | (0.004) | (0.031) | (0.002) |
| Media | −2.430 | −2.783 | −2.625** | −2.225 | −2.478 | −2.400** | −2.519 | −3.147 | −2.828** | −2.815* | −2.750 | −3.116*** |
| | (0.1628) | (0.246) | (0.016) | (0.212) | (0.330) | (0.024) | (0.145) | (0.177) | (0.011) | (0.098) | (0.261) | (0.007) |
| FMCG | −3.825** | −3.380 | −1.909* | −4.065** | −3.291 | −2.655** | −3.952** | −3.260 | −2.776** | −4.147** | −3.817* | −1.425 |
| | (0.0259) | (0.128) | (0.059) | (0.019) | (0.141) | (0.015) | (0.022) | (0.146) | (0.012) | (0.017) | (0.079) | (0.133) |
| Pharma | −2.591 | −2.404 | −2.541** | −1.653** | −1.012*** | −1.475** | −1.426** | −0.313*** | −1.268** | −1.419* | −0.768*** | −1.286** |
| | (0.1324) | (0.354) | (0.019) | (0.041) | (0.008) | (0.012) | (0.006) | (0.009) | (0.017) | (0.051) | (0.009) | (0.016) |
| Metal | −3.372** | −2.416 | −3.788*** | −3.459** | −3.348 | −3.739*** | −2.694 | −2.358 | −3.391*** | −3.078* | −2.852 | −3.336*** |
| | (0.0467) | (0.348) | (0.002) | (0.047) | (0.146) | (0.003) | (0.121) | (0.367) | (0.004) | (0.069) | (0.230) | (0.005) |
| Realty | −3.009* | −2.910 | −3.267*** | −2.711 | −2.366 | −3.177*** | −5.230*** | −5.450** | −5.567*** | −5.448*** | −6.084*** | −5.857*** |
| | (0.0811) | (0.229) | (0.006) | (0.118) | (0.365) | (0.006) | (0.005) | (0.015) | (0.000) | (0.004) | (0.008) | (0.000) |
| IT | −3.491** | −2.931 | −2.100** | −3.233* | −2.527 | −2.112** | −2.477** | −2.235** | −1.242** | −2.285** | −2.013* | −1.293** |
| | (0.040) | (0.214) | (0.041) | (0.062) | (0.318) | (0.041) | (0.015) | (0.041) | (0.018) | (0.019) | (0.051) | (0.016) |
| Energy | −4.621*** | −4.130* | −3.692*** | −3.597** | −3.263 | −2.650** | −2.778 | −2.639 | −1.844* | −2.874* | −2.731 | −2.278** |
| | (0.0096) | (0.057) | (0.003) | (0.035) | (0.145) | (0.015) | (0.103) | (0.283) | (0.065) | (0.091) | (0.258) | (0.030) |
| Bank | −5.432*** | −5.682** | −3.912*** | −5.337*** | −5.512** | −1.098 | −6.620*** | −6.352*** | −1.030 | −4.213** | −4.784** | −4.416*** |
| | (0.0038) | (0.012) | (0.002) | (0.004) | (0.014) | (0.221) | (0.001) | (0.007) | (0.244) | (0.019) | (0.037) | (0.001) |

**Notes:** *Shows significant results at the 10% level of significance, ** shows significant results at the 5% level of significance and *** shows significant results at the 1% level of significance.

## *Correlation Results*

Table 5 shows the correlation matrix amongst four variables (two groups) taken for investigating the financial year and calendar year effects respectively. It may be observed that the correlation between March–April and December–January has been strong. The *P* values indicate that these relationships have been significant at the 1% level of significance ($H_{0a}$: *There is no correlation between March and April returns;* $H_{0b}$: *there is no correlation between December and January returns*). Thus, their connection may be further explored with t-test and other models.

## *t-Test Results*

The independent sample *t*-test results have been depicted in Table 6.

The significant difference between two groups of variables, namely March–April and December–January, has been explored with this test. The results depict that null hypotheses ($H_{0a}$: *there is no significant difference between March and April returns;* $H_{0b}$: *there is no significant difference between December and January returns*) may not be rejected, with the corresponding *P* values being >.05 in the case of all sectoral indices. However, as the independent *t*-test remains a test of a very small order, the relationship between variables has been further tested with the OLS model.

**Table 5.** Correlation Matrix.

| Sector | March–April | December–January |
|---|---|---|
| Auto | 0.969*** | 0.879*** |
| | (0.000) | (0.002) |
| FS | 0.967*** | 0.890*** |
| | (0.000) | (0.001) |
| Media | 0.977*** | 0.918*** |
| | (0.000) | (0.000) |
| FMCG | 0.813*** | 0.887*** |
| | (0.008) | (0.001) |
| Pharma | 0.931*** | 0.973*** |
| | (0.000) | (0.000) |
| Metal | 0.987*** | 0.911*** |
| | (0.000) | (0.001) |
| Realty | 0.940*** | 0.889*** |
| | (0.000) | (0.001) |
| IT | 0.901*** | 0.909*** |
| | (0.001) | (0.001) |
| Energy | 0.954*** | 0.858*** |
| | (0.000) | (0.003) |
| Bank | 0.970*** | 0.901*** |
| | (0.000) | (0.001) |

**Notes:** *** Shows significant results at the 1% level of significance.

**Table 6.** Independent Sample *t*-Test.

| Sector | March–April | December–January |
|---|---|---|
| Auto | –0.016 | 0.310 |
| | (0.988) | (0.759) |
| FS | –0.061 | 0.224 |
| | (0.952) | (0.825) |
| Media | –0.012 | 0.220 |
| | (0.991) | (0.828) |
| FMCG | 0.492 | 0.208 |
| | (0.628) | (0.838) |
| Pharma | –0.124 | 0.278 |
| | (0.903) | (0.784) |
| Metal | 0.018 | 0.237 |
| | (0.986) | (0.815) |
| Realty | –0.209 | 0.409 |
| | (0.837) | (0.687) |
| IT | 0.196 | 0.041 |
| | (0.846) | (0.969) |
| Energy | 0.181 | 0.517 |
| | (0.858) | (0.611) |
| Bank | –0.051 | 0.232 |
| | (0.961) | (0.819) |

**Notes:** March–April shows *t*-statistic and corresponding *P* values between March and April returns. Similarly, December–January shows *t*-statistic and corresponding *P* values between December and January returns.

## *Ordinary Least Squares (Financial Year Cycle)*

The regression analysis for the financial year effect (March–April) has been presented in Table 7.

The *P* values for respective coefficients from sectoral indices have been >.05 in all cases except financial services ($H_0$: *There is no significant impact of March returns on April returns*). Thereby, results indicate that March returns have impacted April returns significantly in the case of financial service index returns only. All other sectoral indices have experienced similar outcomes meaning that announcements made in March in annual reports have not shown any significant influence on the next month's returns. The *F*-statistic values have also been significant at the 5% level of significance, indicating that the model is fitted well. In the Jarque–Bera residual diagnostics, most of the probability values have been >.05 ($H_0$: *Residuals are normally distributed*) meaning that residuals are normally distributed. Breusch–Godfrey statistics suggests that the *P* values have been >.05 except in few cases ($H_0$: *There is no autocorrelation between residuals of the model*). Therefore, the residuals do not have autocorrelation issues. Further, Breusch–Pagan statistics indicates that the *P* values have been >.05 except in a rare case ($H_0$: *There is no heteroskedasticity in residuals of the model*). The results indicate that from 2011 to 2023, the financial year effect persisted only for the financial services

**Table 7.** OLS Results (Financial Year Effect).

| Sector | Coefficient | Adjusted *R*-Squared | *F*-Stat | Normality | Autocorrelation | Heteroskedasticity |
|---|---|---|---|---|---|---|
| Auto | 0.029 | 0.923 | 120.525*** | 7.890** | 1.154 | 0.878 |
| | (0.217) | | (0.000) | (0.019) | (0.282) | (0.348) |
| FS | 0.040* | 0.908 | 100.222*** | 0.497 | 3.201* | 0.037 |
| | (0.059) | | (0.000) | (0.779) | (0.073) | (0.846) |
| Media | 0.008 | 0.951 | 193.861*** | 0.492 | 0.314 | 0.073 |
| | (0.746) | | (0.000) | (0.781) | (0.575) | (0.787) |
| FMCG | 0.031 | 0.602 | 16.097*** | 0.555 | 2.102 | 0.967 |
| | (0.340) | | (0.003) | (0.757) | (0.147) | (0.325) |
| Pharma | 0.032 | 0.832 | 50.567*** | 0.311 | 0.053 | 1.106 |
| | (0.228) | | (0.000) | (0.855) | (0.818) | (0.292) |
| Metal | 0.014 | 0.946 | 174.940*** | 0.287 | 4.409** | 4.271** |
| | (0.644) | | (0.000) | (0.866) | (0.035) | (0.038) |
| Realty | 0.026 | 0.854 | 59.282*** | 1.193 | 0.007 | 1.253 |
| | (0.463) | | (0.000) | (0.550) | (0.935) | (0.262) |
| IT | 0.027 | 0.795 | 39.752*** | 1.331 | 4.211** | 0.147 |
| | (0.498) | | (0.000) | (0.513) | (0.040) | (0.701) |
| Energy | 0.022 | 0.845 | 55.387*** | 0.465 | 6.890*** | 1.676 |
| | (0.447) | | (0.000) | (0.792) | (0.008) | (0.195) |
| Bank | 0.034 | 0.924 | 122.719*** | 0.292 | 3.490* | 0.212 |
| | (0.113) | | (0.000) | (0.863) | (0.061) | (0.645) |

**Notes:** *Shows significant results at the 10% level of significance, ** shows significant results at the 5% level of significance and *** shows significant results at the 1% level of significance.

sector. It may be because this sector connects the entire economy, and reaction in this sector takes time to fade away from the following month's returns.

## Ordinary Least Squares (Calendar Year Cycle)

The calendar year effect (December–January) has been examined with the OLS model and presented in Table 8.

The sector-wise analysis reflected that December returns do not have a significant impact on January returns. Thus, the null hypothesis has been failed to be rejected, with *P* values >.05 in all the sectors. ($H_0$: *There is no significant impact of December returns on January returns*.) It may be stated that follow-up reactions in January have not persisted during the period 2011–2023, indicating that national and international shocks have changed the market dynamics. Anomalies may cease to exist due to transparent systems in markets as well as a cautious approach of investors due to the volatile environment in markets and the recent and ongoing events. *F*-statistic from the model has shown all *P* values <.05. Hence, the model has been fitted well. The residual diagnostics of the model have shown that residuals have been normally distributed (Jarque–Bera *P* values > .05; $H_0$: *residuals are normally distributed*). The autocorrelation *P* values are also >.05. ($H_0$: *There is no autocorrelation among the residuals*.) However, heteroskedasticity *P* values have been >.05. ($H_0$: *There is no heteroskedasticity in residuals*.) Few cases of residual issues have been there, but in most of the cases, such issues were not present.

## Discussion of Results

In the present study, a strong association between March–April and December–January effects has been observed. This result is in line with the study conducted by Singh and Das (2020), Selvakumar (2011) and Gouider et al. (2015). Further, Ferrouhi et al. (2021) found the January and December effects in Botswana in African stock markets. High returns were observed in December and March in the Karachi Stock Exchange and in the Pakistan Stock Exchange respectively (Anjum, 2020). As explained by Shen et al. (2020), the causes behind the January effect were the minimum capital gain and restricted supply of stock, and the desire of investors to sell at premium led to minimise the pressure to sell in January.

In addition, returns of the April month have been affected by March returns. Many researchers found different month's effect on the Indian stock market, for example, Sen (2014) and Yadav and Singh (2018) found positive return in the months of September and November. In India, the announcement of the interim Union budget is mostly made during February–April, and the closing of the financial year also affects the March–April months. So, the presence of the financial year effect has been significant in the Indian stock market (Saxena et al., 2021; Singh & Das, 2020) as reported in the present study. In contrast, no month anomalies were there in the Indian stock exchange (Ansari, 2020; Bhatia, 2021;

**Table 8.** Ordinary Least-Squares Results (Calendar Year Effect).

| Sector | Coefficient | Adjusted *R*–Squared | *F*–Stat | Normality | Autocorrelation | Heteroskedasticity |
|---|---|---|---|---|---|---|
| *Auto* | –0.010 | 0.750 | 30.960*** | 0.771 | 1.826 | 0.779 |
| | (0.815) | | (0.000) | (0.680) | (0.176) | (0.377) |
| *FS* | –0.040 | 0.754 | 31.614*** | 0.073 | 0.984 | 1.391 |
| | (0.411) | | (0.000) | (0.964) | (0.321) | (0.238) |
| *Media* | –0.010 | 0.847 | 56.259*** | 0.091 | 2.541 | 6.070** |
| | (0.760) | | (0.000) | (0.955) | (0.110) | (0.013) |
| *FMCG* | 0.006 | 0.760 | 32.725*** | 0.103 | 6.627** | 3.049* |
| | (0.842) | | (0.000) | (0.949) | (0.010) | (0.080) |
| *Pharma* | 0.001 | 0.937 | 150.146*** | 0.239 | 0.879 | 2.614 |
| | (0.973) | | (0.000) | (0.887) | (0.348) | (0.105) |
| *Metal* | –0.035 | 0.860 | 62.580*** | 0.408 | 0.144 | 1.391 |
| | (0.413) | | (0.000) | (0.815) | (0.704) | (0.238) |
| *Realty* | –0.030 | 0.794 | 39.490*** | 0.846 | 4.479** | 0.073 |
| | (0.628) | | (0.000) | (0.654) | (0.034) | (0.787) |
| *IT* | 0.034 | 0.845 | 55.544*** | 0.485 | 1.370 | 0.683 |
| | (0.258) | | (0.000) | (0.784) | (0.241) | (0.408) |
| *Energy* | –0.052 | 0.670 | 21.282** | 0.545 | 3.111* | 0.385 |
| | (0.270) | | (0.001) | (0.761) | (0.077) | (0.535) |
| *Bank* | –0.042 | 0.753 | 31.559*** | 0.354 | 0.576 | 2.155 |
| | (0.415) | | (0.000) | (0.837) | (0.447) | (0.142) |

**Notes:** *Shows significant results at the 10% level of significance, ** shows significant results at the 5% level of significance and *** shows significant results at the 1% level of significance.

Kaur et al., 2019; Kushwah & Munshi, 2018; Nageswari et al., 2011) nor were there Christmas and new year effects (Jain, 2019). In addition, no calendar and financial year effect was noted on the Baltic stock market (Aleknevičienė et al., 2021) and on stock exchange indexes in France, Germany, Italy and Spain (Rossi & Gunardi, 2018).

In this study, the December year effect has been prevalent in various sectors, especially in the banking sector. Further, Singh and Das (2020) also mentioned anomalies in the IT and banking sectors of India. Different sectors show different month effects, such as February and August in the NIFTY Bank index; April in the NIFTY Auto index and March in Nifty Financial Services (Saxena et al., 2021). So, the investors can forecast the trend of prices and earn return in different indices of Indian stock markets.

## Conclusion and Stakeholders' Implications

The present study has emphasised on the analysis of two types of prevalent anomalies for the sectoral indices in Indian markets. The results from the independent sample *t*-test and OLS model have indicated the absence of anomalies during 2011–2023 owing to the absorption of all information in the market prices. It may be concluded from these results that the opportunities from such anomalies have been absent for all indices meaning that information from March and December months may not result in follow-up behaviour in next months. Financial services index has shown rare possibility to explore abnormal returns due to the presence of anomalies in this sector. It may also be added that owing to increasing efficiency of Indian markets in terms of information may be reflected in stock prices leaving lesser scope for traders to utilise volatile stocks.

While planning the proportions in portfolio, investors may pay attention to these sectors and plan, buy and sell strategies. The regulators may keep a watch on corporate disclosures during the months of March and December for avoidance of concentration on any particular sector in markets. Policymakers and managers may lay special emphasis on the shareholders' wealth during these periods owing to expectations of shareholders from announcements. The absence of anomalies across Indian sectoral indices provides least opportunities for investors and traders to obtain any abnormal returns. Their approach however may be watchful while selecting stocks from these sectors, and events of paramount importance may be studied before investing. A comparative analysis of these particular events may be fruitful in planning investments in a better way.

## Scope for Further Study

The study has investigated the financial year and calendar year effects for Indian sectoral indices during 2011–2023. The investigation for anomalies before and after this crisis can be carried out in further studies. The structural breaks may be checked if any analysis for normal and abnormal period may be compared. The examination of other thematic indices from Indian markets and global emerging

sectoral indices across the globe can also be explored with the results of the present study.

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## ORCID iD

Parul Bhatia https://orcid.org/0000-0002-2687-8474

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